

Lecture 7. Symmetrical distributions

Symmetrical Distributions

Let's consider other estimators for θ in our model.

It follows from the symmetry that

$$\theta = f(F) = \frac{1}{1-2\alpha} \int_{\alpha}^{1-\alpha} F^{-1}(x) dx$$

for every $\alpha \in [0, 1/2)$. Therefore, it's natural to use the estimator

$$\hat{\theta} = \frac{1}{1-2\alpha} \int_{\alpha}^{1-\alpha} \hat{F}_n^{-1}(x) dx = \frac{X_{([\alpha n])}([\alpha n] + 1 - n\alpha) + X_{([\alpha n] + 1)} + \dots + X_{(n - [\alpha n])} + X_{(n - [\alpha n] + 1)}(n\alpha - [\alpha n])}{(1-2\alpha)n}$$

Definition 1. The estimator

$$\bar{X}_{\alpha} = \frac{X_{([\alpha n] + 1)} + \dots + X_{(n - [\alpha n])}}{n - 2[\alpha n]},$$

is called the α -trimmed mean.

The α -trimmed mean is similar to $\hat{\theta}$ above and has the same asymptotical distribution.

Example 1. Let's find the asymptotic breakdown point for $\hat{\theta}$. Obviously, $b(1) = \infty$. Moreover, $b(\alpha + \varepsilon) = \infty$ for every $\varepsilon > 0$ since

$$\frac{1}{1-2\alpha} \int_{\alpha}^{1-\alpha} F_n^{-1}(x) dx = \frac{1}{1-2\alpha} \int_{\alpha+\varepsilon}^{1-\alpha} F^{-1}(x) dx - n\varepsilon \rightarrow \infty, \quad n \rightarrow \infty,$$

for $F_n(x) = F_0(x)I_{x \geq F_0^{-1}(\alpha+\varepsilon)} + (\alpha + \varepsilon)I_{x \in [-n, F_0^{-1}(\alpha+\varepsilon)]}$. On the other hand, $b(\alpha - \varepsilon) < \infty$, since $\forall \varepsilon > 0$ we have

$$F^{-1}(x) \geq F_0^{-1}(x - (\alpha - \varepsilon)) - \alpha + \varepsilon \geq F_0^{-1}(\varepsilon) - 1 > -\infty$$

as $x \in [\alpha, 1 - \alpha]$. Therefore, $\varepsilon^* = \alpha$.

For any $\alpha > 0$ this estimator is robust. On the other hand, it's an asymptotically normal estimator in $\mathcal{F} = \mathcal{F}_{symm}$

Theorem 1. As $n \rightarrow \infty$

$$\sqrt{n}(f(\hat{F}_n) - f(F)) \xrightarrow{d} Z \sim \mathcal{N}(0, \sigma_{\alpha}^2),$$

where

$$\sigma_{\alpha}^2 = \frac{2}{(1-2\alpha)^2} \left(\int_{\theta}^{x_{1-\alpha}} (t - \theta)^2 f(t) dt + \alpha(x_{1-\alpha} - \theta)^2 \right).$$

Доказательство. Let

$$F_n(u) = F(u) + \varepsilon D(u).$$

Then

$$F_n^{-1}(x) = F^{-1}(-\varepsilon D(F_n^{-1}(x))) = F^{-1}(x) - (F^{-1}(x))' \varepsilon D(F_n^{-1}(x)) + o(\varepsilon).$$

Therefore

$$\begin{aligned} \frac{F_n^{-1}(x) - F^{-1}(x)}{\varepsilon} &= -(F^{-1}(x))'(D(F^{-1}(x) + O(\varepsilon))) + o(\varepsilon) = \\ &= -\frac{1}{p(F^{-1}(x))} D(F^{-1}(x)) + o(\varepsilon), \quad \varepsilon \rightarrow 0. \end{aligned}$$

Last equality holds as $F^{-1}(x)$ is a continuity point of D , therefore it holds for a.s. $x \in [\alpha, 1 - \alpha]$. Notice that $o(\varepsilon)$ is uniformly small on $x \in [\alpha, 1 - \alpha]$. Thus,

$$\frac{f(F + \varepsilon D) - f(F)}{\varepsilon} \rightarrow L_F(D; f) = -\frac{1}{1-2\alpha} \int_{\alpha}^{1-\alpha} \frac{1}{p(F^{-1}(u))} D(F^{-1}(u)) du.$$

Therefore,

$$I_F(x) = \frac{1}{1-2\alpha} \int_{\alpha}^{1-\alpha} \frac{u}{p(F^{-1}(u))} du - \frac{1}{1-2\alpha} \int_{\alpha}^{1-\alpha} \frac{1}{p(F^{-1}(u))} \delta_x(F^{-1}(u)) du. \quad (1)$$

The first term in (1) equals to

$$\int_{\alpha}^{1-\alpha} \frac{u}{p(F^{-1}(u))} du = \int_{x_{\alpha}}^{x_{1-\alpha}} \frac{F(v)}{f(v)} dF(v) = \int_{x_{\alpha}}^{x_{1-\alpha}} F(v) dv = \frac{x_{1-\alpha} - x_{\alpha}}{2}$$

since $F(v - \theta) - 1/2$ is odd function. The second term of (1) is

$$\int_{\alpha}^{1-\alpha} \frac{1}{p(F^{-1}(u))} \delta_x(F^{-1}(u)) du = \int_{x_{\alpha}}^{x_{1-\alpha}} \delta_x(v) dv.$$

As $x > x_{1-\alpha}$ function $\delta_x(v)$ equals to 0, as $x < x_{\alpha}$ $\delta_x(v)$ equals to 1, as $x_{\alpha} < x < x_{1-\alpha}$ integral of δ_v equals to $x_{1-\alpha} - x$. Therefore

$$I_F(f) = \begin{cases} -\frac{x_{1-\alpha} - x_{\alpha}}{2(1-2\alpha)}, & x < x_{\alpha}, \\ \frac{x - \theta}{1-2\alpha}, & x \in (x_{\alpha}, x_{1-\alpha}), \\ \frac{x_{1-\alpha} - x_{\alpha}}{2(1-2\alpha)}, & x \geq x_{1-\alpha}, \end{cases} \quad \sigma^2(F) = \frac{2}{(1-2\alpha)^2} \left(\int_{\theta}^{x_{1-\alpha}} (t - \theta)^2 f(t) dt + \alpha(x_{1-\alpha} - \theta)^2 \right).$$

□

Problem 1. Prove that

$$\sigma^2(F) \leq \frac{\sigma^2}{(1-2\alpha)^2},$$

where $\sigma^2 = \mathbf{D}X$ and the inequality becomes an equality for some F .

So, the α -trimmed mean is a good robust alternative to the sample mean. However, as the asymptotic breakdown point of X_{α} equals $1/3$, $\sigma^2(F)$ can be 9 times larger than σ^2 .